

# WHAT IS (PROPENSITY SCORE) MATCHING?

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# (PS)MATCHING IS EXTREMELY POPULAR...

- 270,000 entries by googling: propensity score matching
- 13,000 downloads of `-psmatch2-`  
501<sup>st</sup> of 1,100,000 items in the RePEc/IDEA database
- >1,500 support emails
  - Europe, US, Canada, Central + South America, former SU, Australia, Asia, Africa and the Middle East
  - epidemiology, sociology, economics, statistics, criminology, agricultural economics, health economics, transport economics, public health, nutrition, paediatrics, biostatistics, finance, urban planning, geography and geosciences

## WHAT IS (PS)MATCHING?

(PS)Matching is a method/device to make two groups look the same.

# Roadmap

1. The counterfactual concept of causality
2. What is matching?
3. How do we use it?
4. Should we use it?



# THE COUNTERFACTUAL CONCEPT OF CAUSALITY

## The Evaluation Problem

to evaluate average *causal effects* of a 'treatment' on an outcome.

## The Potential Outcome model

$Y_1$	Outcome under treatment
$Y_0$	Outcome without treatment
$Y_1 - Y_0$	Treatment effect
$D \in \{0, 1\}$	Treatment indicator
$Y = \begin{cases} Y_0 & \text{if } D = 0 \\ Y_1 & \text{if } D = 1 \end{cases}$	Observed outcome
$X$	Set of observed characteristics

## The parameters of interest

- ATT  $\equiv E(Y_1 - Y_0 | D=1) = E(Y | D=1) - E(Y_0 | D=1)$
- ATNT  $\equiv E(Y_1 - Y_0 | D=0) = E(Y_1 | D=0) - E(Y | D=0)$
- ATE  $\equiv E(Y_1 - Y_0) = ATT \cdot P(D=1) + ATNT \cdot P(D=0)$

## The Fundamental Problem of Causal Inference

Need to invoke (untestable) assumptions to identify **average unobserved counterfactuals**.

### MATCHING METHODS – INTUITION (FOR ATT)

*Ex post* mimic a RCT by constructing a suitable comparison group by carefully matching treated and non-treated

→ selected comparison group is as similar as possible to the treatment group...in terms of their *observable* characteristics

# MATCHING METHODS – ASSUMPTIONS

1. Identifying assumption: **Selection on Observables**

(conditional independence CIA, exogeneity, ignorability, unconfoundedness)

All the relevant differences between treated and non-treated are captured in  $X$ :

$$\text{ATT: } E(Y_0 | X, D=1) = E(Y_0 | X, D=0)$$

$$\text{ATNT: } E(Y_1 | X, D=1) = E(Y_1 | X, D=0)$$

ATE: both

2. To give it empirical content: **Common Support**

We observe participants and non-participants with the same characteristics:

$$\text{ATT: } P(D=1 | X) < 1$$

$$\text{ATNT: } 0 < P(D=1 | X)$$

$$\text{ATE: } 0 < P(D=1 | X) < 1$$

⇒ can use the (observed) mean outcome of the non-treated to estimate the mean (counterfactual) outcome the treated would have had they not been treated.

# OPERATIONALISING MATCHING METHODS

## Curse of dimensionality

- impose linearity in the parameters (regression analysis)
- choose a distance metric

### ❖ Mahalanobis metric

$$d(i,j) = (\mathbf{X}_i - \mathbf{X}_j)' V^{-1} (\mathbf{X}_i - \mathbf{X}_j)$$

### ❖ Propensity Score $p(x) \equiv P(D=1 | X=x)$

Conditional treatment probability (given confounders X)

The propensity score is a balancing score, i.e.  
 $X \perp D | p(X)$

If CIA holds given  $X \rightarrow$  CIA holds given  $p(X)$

# Overview of Matching Estimators

1. pair to each treated  $i$  some group of ‘comparable’ non-treated individuals
2. associate to the outcome  $y_i$  of treated  $i$ , a matched outcome  $\hat{y}_i$  given by the (weighted) outcomes of his ‘neighbours’ in the comparison group:

$$\hat{y}_i = \sum_{j \in C^0(p_i)} w_{ij} y_j$$

- $C^0(p_i)$  = set of neighbours of treated  $i$  in the  $D=0$  group
- $w_{ij}$  = weight on non-treated  $j$  in forming a comparison with treated  $i$ , where  $\sum_{j \in C^0(p_i)} w_{ij} = 1$

General form of the matching estimator for ATT (within  $S_{10}$ ):

$$\hat{ATT} = \frac{1}{\#(D=1 \cap S_{10})} \sum_{i \in \{D_i=1 \cap S_{10}\}} \{y_i - \hat{y}_i\}$$

$$= E(Y \mid \text{treated on } S_{10}) - E(Y \mid \text{matched/reweighted non-treated})$$

### **TRADITIONAL MATCHING ESTIMATORS**

One-to-one matching

- with or without replacement
- nearest neighbour or within caliper

### **SIMPLE SMOOTHED MATCHING ESTIMATORS**

- $K$ -nearest neighbours
  - with or without replacement
  - nearest neighbour or within caliper
- Radius matching

### **WEIGHTED SMOOTHED MATCHING ESTIMATORS**

- Kernel-based matching
- Local linear regression-based matching
  - bandwidth choice
  - kernel choice

## Checking matching quality

Check (and possibly improve on) balancing of observables

- for each variable
- overall measures

$$D \perp X \mid \hat{p}(X)$$

## Inference

- naïve variance
- bootstrapping
- Abadie-Imbens heteroskedasticity-robust standard errors when matching on  $X$
- Abadie-Imbens analytical asy std errors taking into account estimation of  $e(X)$  for PS nearest neighbour(s) matching with replacement

# MATCHING VS OLS

- **same identifying assumption**

If unobserved confounders, just as biased as OLS – internal validity

- **avoids any additional assumption**

- (1) **COMMON SUPPORT**

Matching performed only over  $Sup_{10}$ , hence compares only comparable people

Might recover a different causal impact:  $ATT(Sup_{10}) \neq ATT(Sup_1)$  – external validity

- (2) **NON-PARAMETRIC**

Avoids potential misspecification of  $E(Y_0 | X)$

Allows for arbitrary  $X$ -heterogeneity in impacts  $E(Y_1 - Y_0 | X)$

⇒ Matching focuses on **comparability** in terms of **observables**,  
i.e. on constructing a suitable comparison group by carefully matching treated and non-treated on  $X$  / reweighting the non-treated to realign their  $X$

But: if OLS is correctly specified, OLS is more efficient.

**BUT we don't need matching to make OLS less parametric...**

## **FULLY INTERACTED OLS -FILM-**

$$Y = m_0(X_1, X_2) + \delta D + \delta_1(X_1 D) + \delta_2(X_2 D) + \delta_{12}(X_1 X_2 D) + e$$

$$\beta_{ATT} = \delta + \delta_1 \bar{X}_{1|D=1} + \delta_2 \bar{X}_{2|D=1} + \delta_{12} (\overline{X_1 X_2})_{|D=1}$$

$$\beta_{ATNT} = \delta + \delta_1 \bar{X}_{1|D=0} + \delta_2 \bar{X}_{2|D=0} + \delta_{12} (\overline{X_1 X_2})_{|D=0}$$

$$\beta_{ATE} = \delta + \delta_1 \bar{X}_1 + \delta_2 \bar{X}_2 + \delta_{12} (\overline{X_1 X_2})$$

Can F-test for presence of heterogeneous effects.

# STILL, matching ( $\neq$ OLS) highlights comparability of groups

## Check matching quality

- Propensity score
  - more 'structural' model
  - more flexible specification
  - probit/logit
  - probability/index/odds ratio
- Matching
  - metric:  $X$ ,  $\hat{p}(X)$  or  $\{X, \hat{p}(X)\}$
  - type of matching
  - smoothing parameters
  - common support
- Assessment of matching quality

CAN we get the two groups balanced (in terms of  $X$ )?  
[Think back to RCT...]

# STRENGTHS AND WEAKNESSES

## 😊 Advantages 😊

- controls for selection on observables and on observably heterogeneous impacts
- non-(or semi-) parametric:  
no specific form for outcome equation, decision process or either unobservable term
- $Sup_{10}$ : compare only comparable people and help determining which results reliable
- flexible and easy

## 😞 Disadvantages 😞

- selection on observables: matching as good as its  $X$ 's
- restricting to  $Sup_{10}$  may change parameter being estimated → unable to identify ATT
- data hungry

# EXAMPLE: IMPACT OF NSW

Very famous data in the evaluation literature, combining treatment and controls from a randomised evaluation of the NSW Demonstration with non-experimental individuals drawn from various sources.

NSW male treated (297) with  
male comparisons drawn from the PSID (2,490)

$Y$  = real earnings in 1978

$X$  = age, ethnicity (black and hisp), education (years and <12 years), real pre-programme earnings in 1975

## COMPARABILITY OF GROUPS

*NSW trainees vs NSW control group*

Variable	Mean			t-test	
	Treated	Control	%bias	t	p> t
age	24.626	24.447	2.7	0.36	0.721
black	.80135	.8	0.3	0.04	0.965
hispanic	.09428	.11294	-6.1	-0.80	0.422
educ	10.38	10.188	11.2	1.49	0.136
nodegree	.73064	.81412	-20.0	-2.67	0.008
married	.16835	.15765	2.9	0.38	0.701
re75	3066.1	3026.7	0.8	0.10	0.918
Pseudo R2	LR chi2	p>chi2	MeanB	MedB	
0.008	7.83	0.348	6.3	2.9	

True ATT (experimental estimator) = 886\*

NSW trainees vs PSID comparison group

Variable	Mean		%bias	t-test	
	Treated	Control		t	p> t
age	24.626	34.851	-116.6	-16.48	0.000
black	.80135	.2506	132.1	20.86	0.000
hispanic	.09428	.03253	25.5	5.21	0.000
educ	10.38	12.117	-68.6	-9.51	0.000
nodegree	.73064	.30522	94.0	15.10	0.000
married	.16835	.86627	-194.9	-33.02	0.000
re75	3066.1	19063	-156.6	-20.12	0.000
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Pseudo R2	LR chi2	p>chi2	MeanB	MedB	
0.613	1158.40	0.000	112.6	116.6	

→ expect naïve comparison to be downward biased

Naïve estimator = -15,578\*\*\*

# Distribution of $\hat{p}(X)$

## NSW treated

	Percentiles	Smallest
1%	.0072364	.0013841
5%	.0615839	.0023394
10%	.1406408	.0072364
25%	.4338393	.0117305
50%	.728096	
		Largest
75%	.8627535	.9305425
90%	.912396	.9305425
95%	.9244412	.9305425
99%	.9305425	.9402942

## PSID comparisons

	Percentiles	Smallest
1%	1.19e-17	3.36e-68
5%	8.52e-11	1.31e-35
10%	1.29e-08	4.62e-34
25%	5.14e-06	1.00e-29
50%	.0005869	
		Largest
75%	.0184245	.8831188
90%	.1239506	.8924563
95%	.2752407	.9135577
99%	.733402	.9172212

NSW trainees vs **matched** PSID comparison group – nearest neighbour (w/ replac)

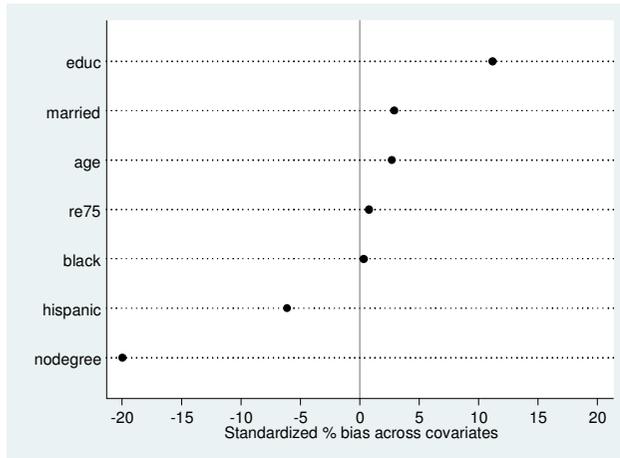
Variable	Mean		%bias	t-test	
	Treated	Control		t	p> t
age	24.626	24.939	-3.6	-0.52	0.606
black	.80135	.79798	0.8	0.10	0.919
hispanic	.09428	.09091	1.4	0.14	0.888
educ	10.38	10.189	7.6	1.05	0.294
nodegree	.73064	.69697	7.4	0.91	0.365
married	.16835	.12795	11.3	1.39	0.166
re75	3066.1	3147.8	-0.8	-0.22	0.823
Pseudo R2	LR chi2	p>chi2	MeanB	MedB	
0.010	7.88	0.343	4.7	3.6	

NSW trainees vs *matched* PSID comparison group – Mahal on X and p(X)

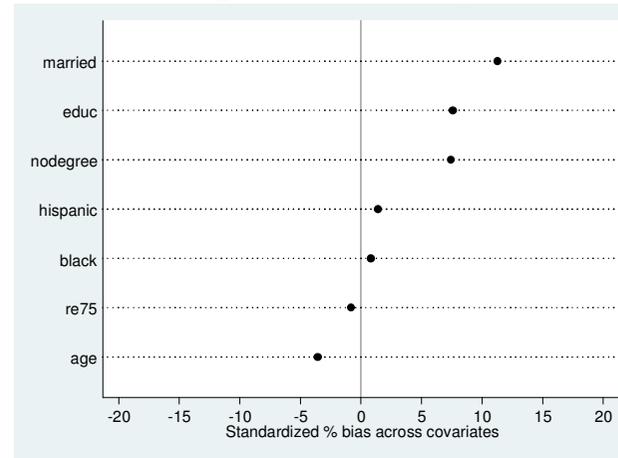
Variable	Mean		%bias	t-test	
	Treated	Control		t	p> t
age	24.626	24.764	-1.6	-0.26	0.792
black	.80135	.80135	0.0	-0.00	1.000
hispanic	.09428	.09428	0.0	0.00	1.000
educ	10.38	10.481	-4.0	-0.69	0.490
nodegree	.73064	.73064	0.0	-0.00	1.000
married	.16835	.17172	-0.9	-0.11	0.913
re75	3066.1	3210.9	-1.4	-0.38	0.705
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Pseudo R2	LR chi2	p>chi2	MeanB	MedB	
0.001	1.16	0.992	1.1	0.9	

# Achieved balancing

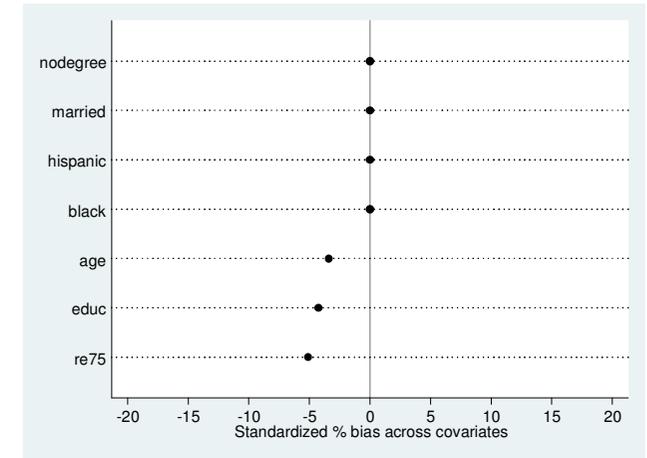
## Randomisation



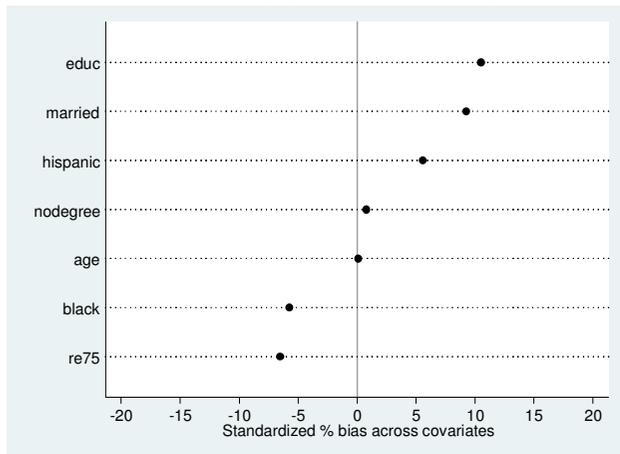
## Nearest Neighbour, with replacement



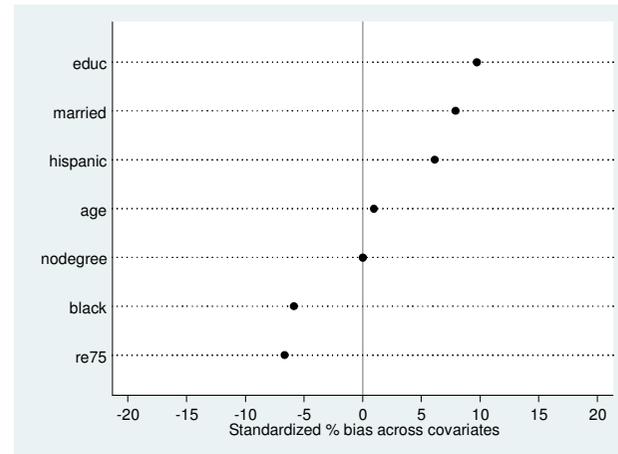
## Mahalanobis



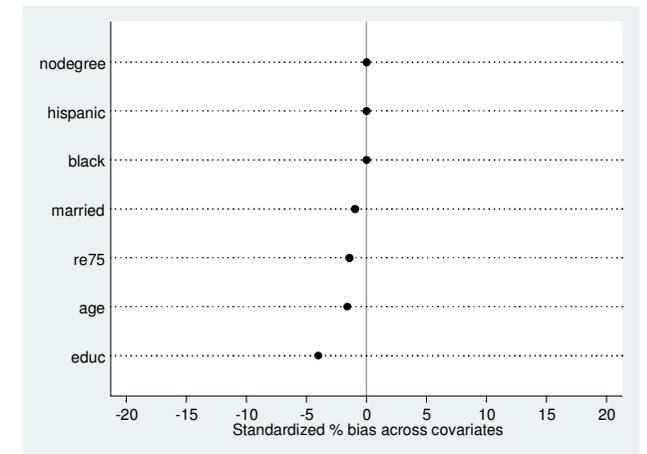
## Kernel (epan, $h=0.01$ )



## Radius ( $r=0.01$ )



## Augmented Mahalanobis



# How many PSID members are we *really* using?

Nearest neighbour (w/ replac)

Kernel (Epan, h=0.01)

psmatch2: weight of matched controls	Freq.
1	73
2	20
3	8
4	3
5	4
6	1
7	2
8	1
10	1
11	2
12	2
19	1
25	1
<b>Total</b>	<b>119</b>

psmatch2: weight of matched controls			
	Percentiles	Smallest	
1%	.0016077	.0016077	
5%	.0016077	.0016077	
10%	.0016077	.0016077	Obs 2488
25%	.0016086	.0016077	Sum of Wgt. 2488
50%	.0017062		Mean .1089228
		Largest	Std. Dev. .6965381
75%	.0241382	8.997245	Variance .4851654
90%	.0779037	10.69014	Skewness 13.42271
95%	.2674301	14.43801	Kurtosis 230.4088
99%	2.285146	15.56199	

## Impact estimates

True ATT (experimental estimator)	886*
Naïve estimator	-15,578***
OLS	-1,458*
FILM	-1,361*
Nearest neighbour (w/ replacement)	551
Kernel (Epan, $h=0.01$ )	-737
Augmented Mahalanobis	-830

# ATNT Average effect of NSW programme had the PSID participated in it

Kernel PS matching (epan,  $h=0.06$ )

```
. psmatch2 treated age black hispanic married educ nodegree re75, out(re78) kernel qui ate
```

Variable	Sample	Treated	Controls	Difference	S.E.	T-stat
re78	Unmatched	5976.35202	21553.9209	-15577.5689	913.328457	-17.06
	ATT	5976.35202	7253.90399	-1277.55197	1878.9332	-0.68
	ATU	21553.9209	8973.94382	-12579.9771	.	.
	ATE			-11375.5206	.	.

Fully interacted regression model

```
. film re78 treated age black hispanic married educ nodegree re75, ate
```

	est.	s.e.	p-value	[95% Conf. Interval]	
OLS	-1457.915	801.6278	0.069	-3029.761	113.9315
FILM					
o att	-1360.8	811.7263	0.094	-2952.449	230.8498
o atu	-12467.76	2542.776	0.000	-17453.69	-7481.834
o ate	-11284.14	2289.46	0.000	-15773.36	-6794.915
F-test of no heterogeneous effects:			F =	6.54	Prob>F = 0.0000

### Nearest neighbour

psmatch2: weight of matched controls	Freq.
1	49
2	22
3	8
4	5
5	2
6	3
7	2
8	1
9	1
10	1
11	2
12	1
13	3
14	1
17	2
19	1
20	1
21	1
25	1
26	1
28	3
31	1
33	1
49	1
53	1
69	1
130	1
159	1
1444	1
<b>Total</b>	<b>119</b>

### Kernel (Epan, $h=0.06$ )

psmatch2: weight of matched controls			
	Percentiles	Smallest	
1%	.0587851	.034638	
5%	.0725001	.0587851	
10%	.0992905	.0587851	Obs 297
25%	.1502837	.0587851	Sum of Wgt. 297
50%	.3663589		Mean 8.383838
		Largest	Std. Dev. 32.23784
75%	.9229564	168.1198	
90%	6.949462	169.0728	Variance 1039.278
95%	25.99836	169.3782	Skewness 4.426013
99%	169.0728	169.7076	Kurtosis 21.0784

ATNT:  $-12,580^{***}$  (matching)  $\approx -12,468^{***}$  (film)

Good the PSID did not go into the programme!

Or is it...? 

And now that we are thinking about it...

Do we really want to know the impact the NSW would have had on the full PSID has they participated?!?

Variable	Unmatched Matched	Mean		%bias	%reduct  bias	t-test	
		Treated	Control			t	p> t
age	Unmatched	34.851	24.626	116.6		16.48	0.000
	Matched	34.851	29.923	56.2	51.8	19.00	0.000
black	Unmatched	.2506	.80135	-132.1		-20.86	0.000
	Matched	.2506	.55964	-74.1	43.9	-23.40	0.000
hispanic	Unmatched	.03253	.09428	-25.5		-5.21	0.000
	Matched	.03253	.01161	8.6	66.1	5.04	0.000
educ	Unmatched	12.117	10.38	68.6		9.51	0.000
	Matched	12.117	10.594	60.2	12.3	20.51	0.000
nodegree	Unmatched	.30522	.73064	-94.0		-15.10	0.000
	Matched	.30522	.54157	-52.2	44.4	-17.38	0.000
married	Unmatched	.86627	.16835	194.9		33.02	0.000
	Matched	.86627	.70206	45.9	76.5	14.37	0.000
re75	Unmatched	19063	3066.1	156.6		20.12	0.000
	Matched	19063	13865	50.9	67.5	15.20	0.000

Sample	Pseudo R2	LR chi2	p>chi2	MeanBias	MedBias
Raw	0.613	1158.40	0.000	112.6	116.6
Matched	0.228	1577.43	0.000	49.7	52.2

# WRAPPING UP...

## SELECTION ON UNOBSERVABLES

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- Set of conditioning  $X$  matters  
⇒ **better data help a lot!**

## SELECTION ON OBSERVABLES

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- Avoid use of functional forms in constructing counterfactual  
⇒ **(matching  $\approx$  fully interacted OLS)  $>$  simple OLS**  
no mis-specification bias
- Compare comparable people  
⇒ **matching  $>$  fully interacted OLS**  
highlight – actual comparability of groups,  
– hence reliability (& relevance) of estimates

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