Running MLwiN from within Stata: the runmlwin command

Modern Modeling Methods (M3) Conference, University of Connecticut 26<sup>th</sup> May 2011

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### INTRODUCTION

## Existing multilevel modelling commands in Stata

- Stata provide the xtmixed, xtmelogit and xtmepoisson commands to fit multilevel models
  - Limited range of models can be specified
  - Computationally quite slow to fit models
- Sophia Rabe-Hesketh and Anders Skrondal provide the gllamm command
  - Wide range of models can be specified
  - Computationally slow to fit models
- Other user-written commands include: hlm, realcomimpute, runmplus, sabre, winbugs

# Multilevel modelling in MLwiN

- 1. Estimation of multilevel models for continuous, binary, ordered categorical, unordered categorical and count data
- 2. Constraints allowing models such as the social relations models and behavioural genetics models to be formulated as multilevel models
- 3. Fast estimation via classical and **Bayesian** methods
- 4. Estimation of multilevel models for cross-classified and multiple membership non-hierarchical data structures
- 5. Estimation of multilevel multivariate response models, multilevel spatial models, multilevel measurement error models, multilevel multiple imputation models and multilevel factor models

# Examples

- 1. Two-level multilevel model
- 2. Growth curve models
- 3. Multilevel models for binary responses
- 4. Simulation studies are easy
- 5. MCMC estimation
- 6. Cross-classified models
- 7. Spatial multilevel models
- 8. Export models to WinBUGS
- 9. Work efficiently
- 10. Resources to help you learn runmlwin

### 1. TWO-LEVEL MULTILEVEL MODELS

# Two-level variance components model

- Inner-London schools exam scores data set
- Classic MLwiN User Manual example
- First analysed by Goldstein et al. (1993)
- Reanalysed by Goldstein (2003), Rabe-Hesketh and Skrondal (2008), Rasbash et al. (2009) and others
- 4059 students nested within 65 schools

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## The runmlwin command syntax

 $normexam_{ij} = \beta_0 + u_j + e_{ij}$  $u_j \sim N(0, \sigma_u^2)$  $e_{ij} \sim N(0, \sigma_e^2)$ 

. runmlwin normexam cons, ///

level2(school: cons) ///

level1(student: cons)

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$\operatorname{hormexam}_{ij} = \beta_{0ij} \operatorname{cons}$	
$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$	
$\begin{bmatrix} u_{0j} \end{bmatrix} \sim \mathbf{N}(0, \ \Omega_u) : \ \Omega_u = \begin{bmatrix} 2 \\ \sigma_u & 0 \end{bmatrix}$	
$\begin{bmatrix} e_{0ij} \end{bmatrix} \sim \mathbf{N}(0, \ \Omega_e) \ : \ \Omega_e = \begin{bmatrix} 2 \\ \sigma_e \ 0 \end{bmatrix}$	
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$\operatorname{normexam}_{ij} \sim N(XB, \Omega)$	
$normexam_{ij} = \beta_{0ij} cons$	
$\beta_{0ij} = -0.013(0.054) + u_{0j} + e_{0ij}$	
$\begin{bmatrix} u_{0j} \end{bmatrix} \sim \mathbf{N}(0, \ \Omega_u) : \ \Omega_u = \begin{bmatrix} 0.169(0.032) \end{bmatrix}$	
$[ \Gamma ] = N(0, \alpha) + \alpha = \Gamma$	
$\left[ e_{0ij} \right]^{\sim \text{IN}(0, \Omega_e)} \cdot \Omega_e^{-} \left[ 0.848(0.019) \right]$	
-2*loglikelihood(IGLS Deviance) = 11010.648(4059 of 4059 cases in use)	
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#### Retrieve the level 2 residuals

 $normexam_{ij} = \beta_0 + u_j + e_{ij}$  $u_j \sim N(0, \sigma_u^2)$  $e_{ij} \sim N(0, \sigma_e^2)$ 

runmlwin normexam cons, ///

level2(school: cons, residuals(u)) ///

level1(student: cons)

## Do not pause in MLwiN

 $normexam_{ij} = \beta_0 + u_j + e_{ij}$  $u_j \sim N(0, \sigma_u^2)$  $e_{ij} \sim N(0, \sigma_e^2)$ 

. runmlwin normexam cons, ///

level2(school: cons, residuals(u)) ///

level1(student: cons) nopause

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//.



- . egen u0rank = rank(u0)
- . serrbar u0 u0se u0rank, scale(1.96) yline(0)



. summarize u0

- . generate u0std = (u0 r(mean))/r(sd)
- . generate u0uniform = (u0rank 0.5)/\_N
- . generate u0nscore = invnorm(u0uniform)

```
. scatter u0std u0nscore, yline(0) xline(0) ///
    ylabel(-3(1)3) xlabel(-3(1)3) aspectratio(1)
```

#### Add covariates

$$\begin{split} normexam_{ij} &= \beta_0 + \beta_1 standlrt_{ij} + \beta_2 girl_{ij} + u_j + e_{ij} \\ & u_j \sim \mathrm{N}(0, \sigma_u^2) \\ & e_{ij} \sim \mathrm{N}(0, \sigma_e^2) \end{split}$$

. runmlwin normexam cons standlrt girl, ///

```
level2(school: cons) ///
```

level1(student: cons) nopause

#### Include a random slope

 $normexam_{ij} = \beta_0 + \beta_1 standlrt_{ij} + \beta_2 girl_{ij} + u_{0j} + u_{1j} standlrt_{ij} + e_{ij}$ 

$$\begin{pmatrix} u_{0j} \\ u_{1j} \end{pmatrix} \sim \mathbb{N} \left\{ \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{u0}^2 \\ \sigma_{u01} & \sigma_{u1}^2 \end{pmatrix} \right\}$$

$$e_{ij} \sim \mathbb{N}(0, \sigma_e^2)$$

. runmlwin normexam cons standlrt girl, ///

level2(school: cons standlrt) ///

level1(student: cons) nopause

# Allow for level 1 heteroskedasticity

 $normexam_{ij} = \beta_0 + \beta_1 standlrt_{ij} + \beta_2 girl_{ij} + u_{0j} + u_{1j} standlrt_{ij}$ 

 $+e_{2ij}girl_{ij}+e_{3ij}boy_{ij}$ 

$$\begin{pmatrix} u_{0j} \\ u_{1j} \end{pmatrix} \sim \mathbf{N} \left\{ \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{u0}^2 \\ \sigma_{u01} & \sigma_{u1}^2 \end{pmatrix} \right\}$$
$$\begin{pmatrix} e_{2ij} \\ e_{3ij} \end{pmatrix} \sim \mathbf{N} \left\{ \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{e2}^2 \\ 0 & \sigma_{e3}^2 \end{pmatrix} \right\}$$

. runmlwin normexam cons standlrt girl, ///

level2(school: cons standlrt) ///
level1(student: girl boy, diagonal) nopause

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Random-effe	cts Parameters	Estimate	e Std. Err.	[95% Conf.	Interval]	
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Level 1:	var(girl) var(boy)	. 525164 . 587434	1 .0152836 5 .0209983	. 4952088 . 5462786	. 5551194 . 6285904	

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#### 2. GROWTH CURVE MODELS

# Child weight data

- Weight gain of Asian children in a British community
- 68 children, one to five measurements per child
- First analysed by Goldstein (1986)
- Re-analysed by Rabe-Hesketh and Skrondal (2008) and others

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• graph twoway ///

(connect weight age, connect(ascending)), ///
ytitle("Weight in Kg") xtitle("Age in years")

#### Growth curve model

$$\begin{split} weight_{ij} &= \beta_0 + \beta_1 age_{ij} + \beta_2 age_{ij}^2 + u_{0j} + u_{1j} age_{ij} + e_{ij} \\ \begin{pmatrix} u_{0j} \\ u_{1j} \end{pmatrix} \sim \mathbb{N} \left\{ \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{u0}^2 \\ \sigma_{u01} & \sigma_{u1}^2 \end{pmatrix} \right\} \\ e_{ij} \sim \mathbb{N}(0, \sigma_e^2) \end{split}$$

. runmlwin weight cons age age2, ///

level2(id: cons age, residuals(u)) ///
level1(occ: cons) nopause

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MLwiN 2.23 mul Normal respons Estimation alg	ltilevel model se model gorithm: <b>IGLS</b>		Number	of obs =	198	<u>-</u>
Level Variabl	No. of Groups	Observat Minimum	ions per Group Average Max	o cimum		
i	id 68	1	2.9	5		
Number of iter Log likelihood Deviance weight	rations = d = -258. = 516 Coef.	7 07785 .1557 Std. Err.	z P> z	[95% Conf.	Interval]	
cons age age2	3.494518 7.704002 -1.660475	.1372489 2 .2394275 3 .0885319 -1	5.460.0002.180.0008.760.000	3.225515 7.234733 -1.833994	3.76352 8.173271 -1.486955	
Random-effe	ects Parameters	Estimate	Std. Err.	[95% Conf.	Interva]]	
Level 2:	var(cons) cov(cons,age) var(age)	.4040045 .088273 .2539857	.1412488 .0812774 .0858503	.1271619 0710279 .0857222	.6808471 .2475738 .4222493	
Level 1:	var(cons)	. 331641	. 0532307	. 227 3107	.4359712	

1.



• generate prediction = ///
\_b[cons]\*cons + \_b[age]\*age + \_b[age2]\*age2 ///
+ u0 + u1\*age

 line prediction age, connect(a) /// ytitle("Predicted weight in Kg")



- label define genderlabel 1 "Boy" 2 "Girl"
- label values gender genderlabel
- graph twoway (line weight age, connect(ascending)), /// by(gender) /// xtitle("Age in years") ytitle("Weight in Kg")

## Growth curve model by gender

 $weight_{ij} = \beta_0 boy_j + \beta_1 boy_j \times age_{ij} + \beta_2 girl_j + \beta_3 girl_j \times age_{ij} + \beta_4 age_{ij}^2$ 

 $+u_{0j}boy_j + u_{1j}boy_j \times age_{ij} + u_{2j}girl_j + u_{3j}girl_j \times age_{ij} + e_{0ij}boy_j + e_{2ij}girl_j$ 

$$\begin{pmatrix} u_{0j} \\ u_{1j} \\ u_{2j} \\ u_{3j} \end{pmatrix} \sim \mathbf{N} \left\{ \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{u0}^2 & & & \\ \sigma_{u01} & \sigma_{u1}^2 & & \\ 0 & 0 & \sigma_{u2}^2 & \\ 0 & 0 & \sigma_{u23} & \sigma_{u3}^2 \end{pmatrix} \right\}$$
$$\begin{pmatrix} e_{0ij} \\ e_{2ij} \end{pmatrix} \sim \mathbf{N} \left\{ \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{e0}^2 & & \\ 0 & \sigma_{e2}^2 \end{pmatrix} \right\}$$

. matrix a = (1, 1, 1, 0, 0, 1, 0, 0, 1, 1)

. runmlwin weight boy boyXage girl girlXage age2, ///

level2(id: boy boyXage girl girlXage, elements(a)) ///
level1(occ: boy girl, diagonal) nopause
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	_	No. of	Obser	vatio	ns per	' Group	).			
Level Variab	le	Groups	Minimum	Av	erage	Мах	cimum			
	id	68	1		2.9		5			
Run time (sec	onds	5) =	1.58							
Number of ite	rati	ions =	7							
Log likelihoo	d	= -24/. - 494	49434 98868							
	1	- 151.	50000							
weight		Coef.	Std. Err.		z P	?> z	[9	5% Conf.	Interval]	
boy		3.78267	.1563113	24.	20 0	0.000	3.	476305	4.089034	
boyXage		7.728288	.2567359	30.	10 (	0.000	7.	225095	8.231481	
girl		3.266411	.1796806	18.		0.000	2.	914244	3.618579	
girlXage		/.30246/	.2341932		04 ( 13 (	000	_1.	04343/ 70118/	/.9014//	
agez		-1.024/45	.0049195	-13.	15 (		-1.	/ 51104	-1.430300	
Random-eff	ects	Parameters	Estim	ate	Std.	Err.	[9	5% Conf.	Interval]	
Level 2:										
		var(boy)	.1553	577	.1659	9469	1	698922	.4806076	
C (	ov(ł	ooy,boyXage)	.102	065	.1232	2655	1	395309	.3436609	
	``	/ar(boyXage)	. 3869	626	.1692	2804	i	0551/9 547155	./18/458	
Var(giri)			0161	196	0864	1426	- 1	533048	185544	
var(girlXage)			.0799	457	.0608	3557	0	393292	.1992206	
Level 1:										
		var(boy)	.4182	827	. 0929	9099	.2	361826	.6003828	
		var(girl)	.2429	176	.0555	5108	.1	341183	.3517168	-
·			•							

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### 3. MULTILEVEL MODELS FOR BINARY RESPONSES

# Guatemalan immunization campaign

- Child immunization data
- 2159 children within 1595 mothers within 161 communities
- First analysed by Pebley, Goldman and Rodriguez (1996) and Rodriguez and Goldman (2001)
- Reanalysed by Rabe-Hesketh and Skrondal (2008) and others

### Three-level binary response model

 $immun_{ijk} \sim \text{Binomial}(1, \pi_{ijk})$ 

 $logit(\pi_{ijk}) = \beta_0 + \beta_1 kid2p_{ijk} + \beta_2 rural_k + \beta_3 pcInd81_k + v_k + u_{jk}$  $v_k \sim N(0, \sigma_u^2)$  $u_{jk} \sim N(0, \sigma_u^2)$ 

. runmlwin immun cons kid2p rural pcInd81, ///

level3(cluster: cons) ///

level2(mom: cons) ///

level1(kid:) ///

discrete(dist(binomial) link(logit) denom(cons)) ///

nopause

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MLwiN 2.23 multi Binomial logit r Estimation algor	ilevel model response model rithm: <b>IGLS, M</b> C	QL1	Number	of obs =	= <b>2159</b>	
Level Variable	No. of Groups	Observa Minimum	ations per Group Average Max	imum		
cluster mom	161 1595	1 1	13.4 1.4	55 3		
Run time (second Number of iterat immun	ds) = 2 tions = Coef. St	2.89 5 td. Err.	z P> z	[95% Conf.	Interval]	
cons kid2p rural pcInd81	1433676 .1 .9173057 .1 5668908 .1 8460267 .1	L721252 L179051 L480174 L797028	-0.830.4057.780.000-3.830.000-4.710.000	4807268 .6862159 8569995 -1.198238	.1939915 1.148395 276782 4938157	
			_			
Random-effect	ts Parameters	Estimat	te Std.Err.	[95% Conf.	Interval]	
Level 3:	var(cons)	. 29608	.0772351	.1447038	. 4474597	
Level 2:	var(cons)	. 36745	.1286113	.1153784	.6195254	

1.

### Refit the model using PQL2

*immunized*<sub>*ijk*</sub>~Binomial(1, $\pi_{ijk}$ )

 $logit(\pi_{ijk}) = \beta_0 + \beta_1 kid2p_{ijk} + \beta_2 rural_k + \beta_3 pcInd81_k + v_k + u_{jk}$  $v_k \sim N(0, \sigma_u^2)$  $u_{ik} \sim N(0, \sigma_u^2)$ 

. runmlwin immun cons kid2p rural pcInd81, ///

level3(cluster: cons) ///

level2(mom: cons) ///

level1(kid:) ///

discrete(d(binomial) 1(logit) de(cons) pql2) ///

initsprevious maxiterations(40) nopause

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Model fitted us	sing initial va	lues speci	fied as paramet	er estimates f	rom previous mo	del 🔺	
MLwiN 2.23 mult Binomial logit Estimation algo	tilevel model response model orithm: IGLS, P	QL2	Number	of obs =	2159		
Level Variable	No. of Groups	Observ Minimum	ations per Grou Average Ma	p ximum			
cluster         161         1         13.4         55           mom         1595         1         1.4         3							
Run time (seco Number of itera	nds) = ations =	8.71 29					
immun	Coef. S	td. Err.	z P> z	[95% Conf.	Interval]		
cons kid2p rural pcInd81	1897211 . 1.363391 . 8506848 . -1.313231 .	2437707 1542231 2157692 2638969	-0.780.4368.840.000-3.940.000-4.980.000	6675029 1.061119 -1.273585 -1.83046	.2880607 1.665663 427785 7960028		
Random-effe	cts Parameters	Estima	te Std. Err.	[95% Conf.	Interval]		
Level 3:	var(cons)	.67231	.86 .1658299	. 3472979	. 997 3393		
Level 2:	var(cons)	2.3526	.2589447	1.845162	2.860206	•	

//.

### 4. SIMULATION STUDIES ARE EASY

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📝 Do-file Editor - rodriguez and goldman (1995).do
```

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```
rodriguez and goldman (1995).do
```

```
1
       set seed 12345
 2
       postfile MQL1 ix fx cx sigmaf sigmac using "MQL1.dta", replace
 3
       set obs 2
 4
       generate cx = n - 1
 5
       expand 10
 6
       sort cx
 7
       generate cid = n
 8
       expand 2
 9
       bysort cid: gen fx = n - 1
10
       expand 10
11
       bysort cid (fx): generate fid = n
12
       expand 2
13
       bysort cid fid: gen ix = n - 1
14
       expand 10
15
       bysort cid fid (ix): gen iid = n
16
       generate cons = 1
17
     - forvalues iteration = 1/10 {
           display n(5) as txt "Iteration " as res "`iteration'" as txt " of " as res "100"
18
19
           generate c = rnormal(0, 1)
20
          bysort cid (fid iid): replace c = c[1]
21
           generate f = rnormal(0,1)
22
          bysort cid fid (iid): replace f = f[1]
23
           generate y = rbinomial(1, invlogit(0*cons + 1*ix + 1*fx + 1*cx + f + c))
24
           runmlwin y cons ix fx cx, level3(cid: cons) level2(fid: cons) level1(iid:) ///
25
               discrete(distribution(binomial) link(logit) denominator(cons)) ///
26
               nopause
27
           post MQL1 ([FP1]ix) ([FP1]fx) ([FP1]cx) (sqrt([RP2]var(cons))) (sqrt([RP3]var(cons)))
28
           drop c f y
29
      Ľι
30
      postclose MQL1
31
       use "MQL1.dta", clear
32
       tabstat ix fx cx sigmaf sigmac, format(%3.2f)
```

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### **5. MCMC ESTIMATION**

### Refit the model using MCMC

 $immun_{ijk} \sim \text{Binomial}(1, \pi_{ijk})$ 

 $logit(\pi_{ijk}) = \beta_0 + \beta_1 kid2p_{ijk} + \beta_2 rural_k + \beta_3 pcInd81_k + v_k + u_{jk}$  $v_k \sim N(0, \sigma_u^2)$  $u_{ik} \sim N(0, \sigma_u^2)$ 

. runmlwin immun cons kid2p rural pcInd81, ///

level3(cluster: cons) ///

level2(mom: cons) ///

level1(kid:) ///

discrete(d(binomial) 1(logit) de(cons) pql2) ///

mcmc(on) initsprevious nopause

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No. of e Groups	Observa Minimum	ations per Average	Group Maximu	Im	
cluster 161 mom 1595		13.4 1.4	5	55 3	
= = = = = = = = = = = = = = = = = = =	500 5000 30 1619.22 866.88 752.34 2371.56				
Mean S	Std. Dev.	z	ESS	[95% Cred.	Interval]
2200957 1.754326 -1.145384 -1.709476	. 3358961 . 2300542 . 2759653 . 3672927	-0.66 7.63 -4.15 -4.65	42 53 56 61 -	889235 1.3257 -1.6628 -2.520294	.5050995 2.208608 6193233 -1.065675
	Maara	Ctd Da		F05% c	und Tutl
cts Parameters	меап	Sta. De	V. ESS	[93% C	rea. Intj
var(cons)	1.161717	. 364123	4 81	.6004681	2.046271
var(cons)	5.934905	5 1.22137	5 21	3. <mark>8</mark> 05253	8.552722
	e No. of Groups r 161 1595 abar) = abar) = of pars (pd) = Mean 2 Mean 2 2200957 1.754326 -1.145384 -1.709476 cts Parameters var(cons) var(cons)	No. of Groups         Observa Minimum           er m         161 1595         1           er m         161 1595         1           er m         161 1595         1           er m         1619.22         30           er mds)         =         1619.22           eabar)         =         866.88           of pars (pd)         =         752.34           =         2371.56           Mean         Std. Dev.          2200957         .3358961           1.754326         .2300542           -1.145384         .2759653           -1.709476         .3672927           ects Parameters         Mean           var(cons)         1.161717           var(cons)         5.934905	No. of Groups         Observations per Minimum           ir m         161 1595         1         13.4 1.4           ir m         161 1595         1         1.4           ir m         1595         1         1.4           ir m         1595         1         1.4           ir m         1619.22         30         30           ir mds)         =         1619.22         30           ir abar)         =         866.88         30           of pars (pd)         =         752.34         3271.56           Mean         Std. Dev.         z          2200957         .3358961         -0.66           1.754326         .2300542         7.63           -1.709476         .3672927         -4.65           icts Parameters         Mean         Std. Dev           var(cons)         1.161717         .364123           var(cons)         5.934905         1.22137	No. of Groups         Observations per Minimum         Group Average         Group Maximu           Inim         161         1         13.4         5           Imm         1595         1         1.4         5           Imm         1619.22         30         30         1         1.4           Imm         Iffiliation         1619.22         30         1         1.5           abar)         =         866.88         30         1         1.5         1           Mean         Std. Dev.         z         ESS         1.754326         .2300542         7.63         53           -1.709476         .3672927         -4.65         61         -           ects Parameters         Mean         Std. Dev.         ESS         Var(cons)         1.161717         .3641234         81           var(cons)         5.934905         1.221375         21	No. of Groups         Observations per Minimum         Group Average         Maximum Maximum           Infinition         1         13.4         55           Infinition         1         13.4         55           Infinition         1         13.4         55           Infinition         1         1.4         3           Infinition         1         1.14         3           Infinition         1         1.14         3           Infinition         1         1.5         1.3257           Infinition         .3672927         -4.65         61         -2.520294           Infinition

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. mcmcsum, trajectories



. mcmcsum [RP2]var(cons), fiveplot

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. mcmcsum [RF	2]var(stand	ilrt)				<b>_</b>		
		[RP2	]var(standlı	rt)				
Mean MCSE of Mean Std. Dev. Mode	.0862781 .0024099 .0467082 .0631075	P 0.5% 2.5% 5% 25%	ercentiles .0193246 .0243268 .0298808 .0520173	Thinned Chain Length Effective Sample Size Raftery Lewis (2.5%) Raftery Lewis (97.5%) Brooks Draper (mean)	5000 99 25770 23976 446390			
		50% 75% 95% 97.5% 99.5%	.0765091 .1100566 .179421 .2023509 .2549108					
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### Run the model for longer

*immunized*<sub>*ijk*</sub>~Binomial(1, $\pi_{ijk}$ )

 $logit(\pi_{ijk}) = \beta_0 + \beta_1 kid2p_{ijk} + \beta_2 rural_k + \beta_3 pcInd81_k + v_k + u_{jk}$  $v_k \sim N(0, \sigma_u^2)$  $u_{ik} \sim N(0, \sigma_u^2)$ 

. runmlwin immunized cons kid2p rural pcInd81, ///

```
level3(cluster: cons) ///
level2(mom: cons) ///
level1(kid:) ///
discrete(d(binomial) l(logit) de(cons) pql2) ///
mcmc(burnin(5000) chain(50000) thinning(10)) ///
initsprevious nopause
```

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Level Variab	No. of Groups	Observa Minimum	tions per Average	r Group Maximu	um	
clust	cluster 161 mom 1595		13.4 1.4	!	55 3	
Burnin Chain Run time (seco Deviance (dbar Deviance (the Effective no. Bayesian DIC	= onds) = r) = tabar) = of pars (pd) = =	5000 50000 257 1641.35 895.18 746.17 2387.51				
immun	Mean	Std. Dev.	z	ESS	[95% Cred.	Interval]
cons kid2p rural pcInd81	2421288 1.730911 -1.089141 -1.681882	. 3079326 . 2184296 . 2954561 . 369106	-0.79 7.92 -3.69 -4.56	424 434 471 587	8276453 1.335295 -1.687306 -2.450633	.3486197 2.19006 5090104 9707532
Random-eff	ects Parameters	s Mean	Std. De	ev. ESS	[95% C	red. Int]
Level 3:	var(cons)	1.127521	. 358036	54 572	. 557 5184	1.94313
Level 2:	var(cons)	5.628252	2 1.2581	13 186	3.587712	8.415654

-

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### 6. CROSS-CLASSIFIED MODELS

# Scottish neighbourhood study on child educational attainment

- Scottish neighbourhood study on child educational attainment
- 2310 students nested within 17 schools and 524 neighbourhoods
- First analysed by Garner and Raudenbush (1991)
- Re-analysed by Rabe-Hesketh and Skrondal (2008), Raudenbush (1993), Raudenbush and Bryk (2002) and others

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> ,793,800	)												
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neighid	2	8	10	15 s	16	17	18	19	20				
26									5				
27							1		8				
30 31							1		2 1				
32							1 2		5				
35							2		3				
37									1				
251						4	1	1	4				
252 253					1	3			1				
256				5		-			2				
259				6	1		2						
260				4			3						
262				5 14		1	1 1						
793 794	1	1	7										
795		1	1										
790	4	1											
798	9	1	1										-
Culture and all 152 Decay	mente											CAD NU	MOVE

### Two-way cross-classified model

$$\begin{split} attain_{i} &= \beta_{0} + u_{\text{schid}(i)}^{(3)} + u_{\text{neighid}(i)}^{(2)} + e_{i} \\ u_{j}^{(3)} &\sim \mathrm{N}\big(0, \sigma_{u(3)}^{2}\big), \qquad u_{j}^{(2)} &\sim \mathrm{N}\big(0, \sigma_{u(2)}^{2}\big), \qquad e_{i} &\sim \mathrm{N}(0, \sigma_{e}^{2}) \end{split}$$

- matrix b = (0, .33, .33, .33)
- . runmlwin attain cons, ///

```
level3(schid: cons) ///
level2(neighid: cons) ///
level1(studentid: cons) ///
```

```
mcmc(cc) initsb(b)
```

5tata/MP 11.2 -	http://www.stata-pre	ss.com/data/mln	nus2/neighbo	orhood.dta -	[Results]		
<u>File E</u> dit <u>D</u> ata	<u>Graphics</u> <u>Statistics</u>	Jser <u>W</u> indow <u>H</u>	elp				8
🚰 🛃 🖷 l 🗐 🖻	• 🏨 •   🛃 •   🛃	🔒   🔜   🕘 🕴					
. matrix b = (	(0,.33,.33,.33)						<u> </u>
. runmlwin att > id: cons) mo	tain cons, level cmc(cc) initsb(b	3(schid: con ) nopause	в) level2	(neighid:	cons) le	vel1(studer	It
MLwiN 2.23 mul Normal respons Estimation alg	ltilevel model se model gorithm: <b>MCMC</b>		Nu	mber of o	bs =	2310	
Level Variabl	No. of Groups	Observat Minimum	ions per Average	Group Maximum			
schi neighi	id 17 id 524	22 1	135.9 4.4	286 16			
Burnin Chain Run time (seco Deviance (dbar Deviance (thet Effective no. Bayesian DIC	= onds) = :abar) = of pars (pd) = =	500 5000 6.88 6039.42 5818.77 220.65 6260.07					
attain	Mean S	td. Dev.	z	ESS [	95% Cred.	Interval]	
cons	.0962458 .	0651659	1.48	228	0349854	. 2170355	
Random-effe	ects Parameters	Mean	Std. Dev	. ESS	[95% CI	red. Int]	
Level 3:	var (cons)	. 0995634	. 0492998	3 2175	.0387661	. 2261923	
Level 2:	var(cons)	. 1422148	.0217524	458	.1033897	.1892418	
Level 1:	var (cons)	. 8002955	.0260423	3 2583	.7513126	. 8540478	•

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1.

### Two-way cross-classified model

 $\begin{aligned} attain_{i} &= \beta_{0} + \beta_{1}p7vrq_{i} + \beta_{2}p7read_{i} + \beta_{3}dadocc_{i} + \beta_{4}daded_{i} \\ &+ \beta_{5}momed_{i} + \beta_{6}dadunemp_{i} + \beta_{7}male_{i} + \beta_{8}deprive_{i} \end{aligned}$ 

 $+u_{\text{schid}(i)}^{(3)}+u_{\text{neighid}(i)}^{(2)}+e_i$ 

 $u_j^{(3)} \sim N(0, \sigma_{u(3)}^2), \qquad u_j^{(2)} \sim N(0, \sigma_{u(2)}^2), \qquad e_i \sim N(0, \sigma_e^2)$ 

. matrix b = (0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 1, 1)

. runmlwin attain cons p7vrq p7read dadocc daded /// % f(x) = 0

momed dadunemp male deprive, ///
level3(schid: cons) ///
level2(neighid: cons) ///
level1(studentid: cons) mcmc(cc) initsb(b)

👩 Stata/MP 11.2 - http	://www.stata-p	oress.com/data/m	ılmus2/neigh	borhood.dta -	[Results]			
<u>File E</u> dit <u>D</u> ata <u>G</u> rag	phics <u>S</u> tatistics	User Window	Help					8
📂 🔒 🌒 🗐 💌 •	il. •   📝 •   🖪	1 1 1 0 0	3					
<pre>. runmlwin attain &gt; dadocc &gt; level3 &gt; level2 &gt; level1 &gt; mcmc(column)</pre>	<pre>runmlwin attain cons p7vrq p7read ///</pre>							
> nopauso MLwiN 2.23 multi Normal response r Estimation algor	<b>e</b> level model model ithm: <b>MCMC</b>		Ν	Number of c	obs =	2310		
Level Variable	No. of Groups	Observa Minimum	ations per Average	r Group Maximun	1			
schid neighid	17 524	22 1	135.9 4.4	286 16	5			
Burnin Chain Run time (seconds Deviance (dbar) Deviance (thetabs Effective no. of Bayesian DIC	s) ar) pars (pd)	= 500 = 5000 = 26.6 = 4744.77 = 4704.11 = 40.67 = 4785.44						
attain	Mean	Std. Dev.	z	ESS [	95% Cred.	Interval]		
cons p7vrq p7read dadocc daded momed dadunemp female deprive	.0351255 .0275779 .0262253 .0080741 .142757 .0605109 1224487 .0558048 1562503	.0291963 .0022758 .0017897 .0013761 .0411453 .0379741 .0468065 .0280615 .0260965	1.20 12.12 14.65 5.87 3.47 1.59 -2.62 1.99 -5.99	1706 4556 . 4689 . 4680 . 5452 . 4703 4505 4699 . 3705 .	0220368 0231709 0226729 0053839 0615814 0132536 2130983 0015048 207776	.0938575 .03204 .0297626 .0107416 .2230719 .1342922 028468 .1117451 1055711		
		I						<b></b>

#### 7. SPATIAL MULTILEVEL MODELS

### Scottish lip cancer

- County level lip cancer counts between 1975 and 1980
- 56 Scottish counties
- First analysed by Clayton and Kaldor (1987)
- Re-analysed by Breslow and Clayton (1993), Leyland and Goldstein (2001), Rabe-Hesketh and Skrondal (2008) and others

📅 Stata/MP 11.2 - http://www.bristol.ac.uk/cm	ım/media/runmlwin/lips1.dta - [Results]
<u>File E</u> dit <u>D</u> ata <u>G</u> raphics <u>S</u> tatistics <u>U</u> ser	Window Help
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Review	
Command <u>_rc</u> 1 use "http://www.bristol.ac.uk/cm	<pre>(R) /(R) /</pre>
	Serial number: 50110514919 Licensed to: Centre for Multilevel Modelling University of Bristol
	Notes: 1. (/m# option or -set memory-) 500.00 MB allocated to data 2. (/v# option or -set maxvar-) 5000 maximum variables
	running C:\Program Files (x86)\Stata11\sysprofile.do
Variables 2	running C:\Users\g]9158\profile.do
Name Label Type Forma	
area Dyte %9.00	. use "http://www.bristol.ac.uk/cmm/media/runmlwin/lips1.dta", clear
obs byte %9.00	.
exp float %9.00	
perc_aff byte %9.0	
offs float %9.0c	
pcons byte %9.0ç	
denom byte %9.0ç	
neigh1 byte %9.0ç	
neigh2 byte %9.0ç	
neigh3 byte %9.0ç	
neigh4 byte %9.0ç	
neigh5 byte %9.0ç	Command
neigh6 byte %9.0g	
neign / byte %9.0	
neigno byte %9.0(	1 1

C:\Users\gl9158\Documents



- use
   "http://www.bristol.ac.uk/cmm/media/runmlwin/lips1.dta",
   clear
- merge 1:1 area using "scotdb.dta"
- spmap obs using "scotcoord.dta", id(area) /// fcolor(Blues) legend(position(10)) /// clmethod(custom) clbreaks(0 5 10 15 20 40)

### **Over-dispersed Poisson model**

 $obs_i \sim Poisson(\pi_i)$ 

 $log(\pi_i) = offs_i + \beta_0 + \beta_1 perc\_aff_i + u_i$  $u_i \sim N(0, \sigma_u^2)$ 

```
runmlwin obs cons perc_aff, ///
    level2(area: cons) ///
    level1(area:) ///
    discrete(dist(poisson) link(log) offset(offs)) ///
    mcmc(chain(50000)) ///
    initsprevious nopause
```

5tata/MP 11.2 - http:	//www.b	ristol.ac.uk/cmn	n/media/r	unmlwin/lips	1.dta - [Resu	ts]					
<u>File E</u> dit <u>D</u> ata <u>G</u> rap	hics <u>S</u> ta	tistics <u>U</u> ser <u>W</u>	<u>indow H</u> e	lp							8
💕 🔒 🌒 🗐 🖻 • 👔	L •   🛃	-   🛃 📑   🔲	00								
<pre>. runmlwin obs cons perc_aff, level2(area:cons) level1(area:) discrete(distribut &gt; ion(poisson) link(log) offset(offs)) mcmc(chain(50000) refresh(500)) initsprev &gt; ious nopause</pre>											Jt 🔺
MLwiN 2.23 mul Poisson respor Estimation alg	ltilev nse mo goritk	vel model odel hm: MCMC,	MQL1			Numbe	r of (	obs	=	56	
Level Variab	le	No. of Groups	Min	Observa imum	tions pe Average	r Gro Ma	up aximur	n			
are	ea	56		1	1.0		]	1			
Burnin Chain Run time (seco Deviance (dbar Deviance (thet Effective no. Bayesian DIC	onds) r) tabar] of pa	= = = ars (pd) = =	=	500 50000 10.3 270.38 230.65 39.73 310.11							
obs		Mean	Std. I	Dev.	z	ESS	I	[95% Cı	red.	Interval]	
cons perc_aff	4 _(	4835851 0675449	.1628 .0143	447 747	-2.97 4.70	494 475	-	.814369 .038883	93 35	1711187 .0956604	
Random-effects Parameters Mean					Std. D	ev.	ESS	[95	5% Cr	ed. Int]	
Level 2:		var(cons)	)	3852672	.11257	43	6787	. 2107	7703	.645173	

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### CAR model

 $obs_i \sim Poisson(\pi_i)$ 

 $\log(\pi_i) = offs_i + \beta_0 + \beta_1 perc\_aff_i + u_i$ 

$$u_i \sim N\left(\bar{u}_i, \frac{\sigma_u^2}{r_i}\right), \qquad \bar{u}_i = \sum_{\substack{j \in neigh(i)}} \frac{w_{i,j}u_i}{r_i}$$

. runmlwin obs perc\_aff, ///

level2(area: cons, carids(neigh1-neigh11) ///
carweights(wcar1-wcar11)) ///

level1(cons:) ///

discrete(dist(poisson) link(log) offset(offs)) ///

Stata/MP 11.2 - http:	//ww	w.bristol.ac.uk/cmm/i	media/runmlwin/lips	1.dta - [Results	]			
<u>File Edit Data Grap</u>	hics	Statistics User Wind	low <u>H</u> elp					8
	h - I		0.03					
. runmlwin obs > level1(cons > fresh(500)) MLwiN 2.23 mu Poisson respon Estimation alg	s pe s:) ini ltil nse gori	erc_aff, leve discrete(dis itsprevious r evel model model thm: MCMC, M	212(area: co tribution(p lopause QL1	ns, carid oisson) l N	s(neigh1 ink(log) umber of	- <b>neigh11) d offset(off</b> obs =	carweights(w Fs)) mcmc(ch = 56	▲ wcar1-wcar11)) hain(50000) re
Level Variab	le	No. of Groups	Observa Minimum	tions per Average	Group Maxim	um		
are	ea	56	1	1.0		1		
Burnin Chain Run time (seco Deviance (dbai Deviance (the Effective no. Bayesian DIC	onds r) taba of	= ;) = ur) = pars (pd) = =	500 50000 9.67 268.77 240.42 28.35 297.13					
obs		Mean S	td. Dev.	z	ESS	[95% Cred.	Interval]	
perc_aff		.035667 .	0128288	2.78	354	.0090298	.0591634	
Random-effe	ects	Parameters	Mean	Std. De	v. ESS	[95% c	[red. Int]	
Level 2:		var(cons)	. 5337886	. 190098	5 3602	.2512767	<b>. 9866003</b>	

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### **Convolution model**

$$\begin{split} obs_i \sim \operatorname{Poisson}(\pi_i) \\ \log(\pi_i) &= offs_i + \beta_0 + \beta_1 perc_{aff_i} + v_i + u_i \\ v_i \sim \operatorname{N}\left(\bar{v}_i, \frac{\sigma_v^2}{r_i}\right), \qquad \bar{v}_i = \sum_{j \in neigh(i)} \frac{w_{i,j}v_i}{r_i} \\ u_i \sim \operatorname{N}(0, \sigma_u^2) \end{split}$$

. runmlwin observed cons perc\_aff, ///

level3(area: cons, carids(neigh1-neigh11) ///
carweights(wcar1-wcar11)) ///
level2(area: cons) level1(county:) ///
discrete(d(binomial) 1(log) offset(offs)) ////

📅 Stata/MP 11.2 - http://www.bristol.ac.uk/cmm/media/runmlwin/lips1.dta - [Results]								
File Edit Data Graphics	Statistics User Windo	w <u>H</u> elp					8	
MLwiN 2.23 multi Poisson response Estimation algor	level model model ithm: MCMC, MC	QL1	N	umber of o	obs =	56		
Level Variable	No. of Groups	Observa Minimum	tions per Average	Group Maximu	— m			
area	56	1	1.0	:	1			
Burnin Chain Run time (second Deviance (dbar) Deviance (thetab Effective no. of Bayesian DIC	= = = ar) = pars (pd) = =	500 50000 14.6 267.86 238.12 29.74 297.60						
obs	Mean St	d. Dev.	z	ESS	[95% Cred.	Interval]		
cons perc_aff	.8812095 1. .0365398 .0	.762997 )136445	0.50 2.68	26 – 472	2.068313 .0084325	3.722563 .0622706		
Random-effect	s Parameters	Mean	Std. De	v. ESS	[95% C	red. Int]		
Level 3:	var(cons)	. 4903258	. 189458	7 2166	.2036356	.9411708		
Level 2:	var(cons)	. 0236907	.032455	9 208	.0006044	.1154397		

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### 8. EXPORT MODELS TO WINBUGS
### The runmlwin command syntax

 $normexam_{ij} = \beta_0 + \beta_1 standlrt_{ij} + u_j + e_{ij}$  $u_j \sim N(0, \sigma_u^2)$  $e_{ij} \sim N(0, \sigma_e^2)$ 

. runmlwin normexam cons standlrt, ///

```
level2(school: cons) ///
level1(student: cons) ///
mcmc(savewinbugs(model(m.txt) inits(i.txt) ///
data(d.txt) nofit)) ///
```

initsprevious nopause

```
Viewer (#1) [view m.txt]
             🙀 view m.txt
                                                                      P.
             What's New
 Advice.
       Contents
                      News
# WINBUGS 1.4 code generated from MLwiN program
#----MODEL Definition-----
model
# Level 1 definition
for(i in 1:N) {
normexam[i] ~ dnorm(mu[i],tau)
mu[i]<- beta[1] * cons[i]</pre>
+ beta[2] * standlrt[i]
+ u2[schoo][i]] * cons[i]
 }
# Higher level definitions
for (j in 1:n2) {
u2[j] \sim dnorm(0,tau.u2)
# Priors for fixed effects
for (k \text{ in } 1:2) \{ \text{ beta}[k] \sim \text{df}[at()) \}
# Priors for random terms
tau ~ dgamma(0.001000,0.001000)
sigma2 <- 1/tau
tau.u2 ~ dgamma(0.001000,0.001000)
sigma2.u2 <- 1/tau.u2</pre>
```

### t-distributed level 2 residuals

 $\begin{aligned} normexam_{ij} &= \beta_0 + \beta_1 standlrt_{ij} + u_j + e_{ij} \\ u_j \sim t(0, \sigma_u^2, \nu) \\ e_{ij} \sim \mathrm{N}(0, \sigma_e^2) \end{aligned}$ 

```
Viewer (#1) [view m.txt]
             🙀 view m.txt
                                                                     P.
             What's New
 Advice.
       Contents
                     News
# WINBUGS 1.4 code generated from MLwiN program
#----MODEL Definition-----
model
# Level 1 definition
for(i in 1:N) {
normexam[i] ~ dnorm(mu[i],tau)
mu[i]<- beta[1] * cons[i]</pre>
+ beta[2] * standlrt[i]
+ u2[schoo][i]] * cons[i]
 }
# Higher level definitions
for (i in 1:n2) {
u2[j] \sim dt(0,tau.u2,df)
# Priors for fixed effects
for (k \text{ in } 1:2) \{ \text{ beta}[k] \sim \text{df}[at()) \}
# Priors for random terms
tau ~ dgamma(0.001000,0.001000)
sigma2 <- 1/tau
tau.u2 ~ dgamma(0.001000,0.001000)
sigma2.u2 <- 1/tau.u2
df \sim dunif(2,200)
```

## The winbugs suite of commands

. wbscript , ///

model("`c(pwd)'\m.txt") inits("`c(pwd)'\i.txt") ///

data("`c(pwd)'\d.txt") coda("`c(pwd)'\out") ///

set(df) burn(500) update(5000) ///

saving("`c(pwd)'\script.txt", replace) quit

. wbrun, script("`c(pwd)'\script.txt") ///

winbugs("C:\Users\gl9158\WinBUGS14\winbugs14.exe")

```
. wbcoda, root("`c(pwd)'\out") clear
```

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#### File Tools Edit Attributes Info Model Inference Options Doodle Map Text Window Help

WinBUGS Licence		
Licence Agreement	ے ا تقالم ا	×
LICENCE AGREEMENT LICENCE AGREE	display(log) check(Q:/C-modelling/runmlwin/presentations/2011-05-26 Connecticut/m_modified.txt) model is syntactically correct data(Q:/C-modelling/runmlwin/presentations/2011-05-26 Connecticut/d.txt) data loaded compile(1) model compiled inits(1,Q:/C-modelling/runmlwin/presentations/2011-05-26 Connecticut/i_modified.txt) model is initialized gen.inits() command #Bugs:gen.inits cannot be executed (is greyed out) update(500) set(df) update(5000) coda(*,Q:/C-modelling/runmlwin/presentations/2011-05-26 Connecticut/out)	
		-

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. mcmcsum df, trajectories variables

🃸 Stata/MP 11.2 - [Re	sults]					
File Edit Data Gra	aphics Statistics Use	er Window H	lelp			8
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. mcmcsum df,	, variables					<b></b>
			df			
Mean MCSE of Mean Std. Dev. Mode	100.3878 4.172605 57.77448 21.98396	0.5% 2.5% 5% 25%	Percentiles 3.102855 7.934025 11.977 49.3425	Thinned Chain Length Effective Sample Size Raftery Lewis (2.5%) Raftery Lewis (97.5%) Brooks Draper (mean)	5000 107 57207 5464 13383	
		50%	102.25	Brooks Braper (meany	19909	
		75% 95% 97.5% 99.5%	149.4 191.2 195.8 199.3			
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:  .						
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### 9. WORK EFFICIENTLY

```
📝 Do-file Editor - Amsterdam2011.do
```

<u>File Edit Tools View</u>

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Amsterdam2011.do

```
41
       * Open the tutorial data set
42
       use "http://www.bristol.ac.uk/cmm/media/runmlwin/tutorial.dta", clear
43
44
45
       * Fit a two-level (students within schools) variance components model to
46
       * a continuous educational response variable, normexam. Note, you will need
47
       * to click the "Resume Macro" button twice in MLwiN to return the model
48
       * results to the Stata output window.
49
       runmlwin normexam cons, ///
50
           level2(school: cons) ///
51
           level1(student: cons)
52
53
       * Generate a boy dummy variable
54
       generate boy = 1 - girl
55
56
       * Extend the previous model to include fixed part covariates, a random school
       * level slope and separate level 1 residuals for boys and girls. The runmlwin
57
58
       * command also requests that runmlwin extracts the predicted values for the
59
       * school level residuals from MLwiN and returns them to Stata. The nopause
60
       * option prevents MLwiN from pausing before and after model estimation and so
61
       * returns the model results automatically to Stata.
62
       runmlwin normexam cons standlrt girl, ///
63
           level2(school: cons standlrt, residuals(u)) ///
64
           level1(student: girl boy, diagonal) nopause
65
66
       * Perform a Wald test to compare the boy and girl residual variances
67
       test [RP1]var(girl) = [RP1]var(boy)
68
       * Preserve the data as we will shortly be collapsing the data to the school
69
       * level, but afterwards we will want to return to the original data
70
71
       preserve
72
73
           * Tag one child in each school
74
           egen pickone = tag(school)
75
76
           * Collapse the data to one row per school
•
```

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Ready

### 10. RESOURCES TO HELP YOU LEARN RUNMLWIN

+ + C = M help runmlwin	Viewer (#1)	) [help runmlwin]	×
	+ +	C 🖶 🏹 help runmlwin	R
Advice Contents What's New News	Advice	Contents What's New News	

help for **runmlwin** 

#### <u>Title</u>

runmlwin - Running the MLwiN multilevel modelling package from within Stata

#### <u>Syntax</u>

runmlwin responses\_and\_fixed\_part, random\_part [discrete(discrete\_options)] [options]

where the syntax of responses\_and\_fixed\_part is one of the following

for univariate continuous, binary, proportion and count response models

depvar indepvars [if] [in]

for univariate multinomial ordered and unordered response models

depvar indepvars1 [(indepvars2, contrast(numlist)) ... ] [if] [in]

where *indepvars1* are those independent variables which appear with separate coefficients in each of every log-odds contrast, while *indepvars2* are those independent variables which appear with common coefficients for those log-odds contrasts specified in **contrast(***numlist***)**. Contrasts can be thought of as the separate "subequations" or "arms" of a multinomial response model. These contrasts are indexed 1,2,... up to the total number of contrasts included in the model. The total number of contrasts will be one less than the number of response categories.

for multivariate response models

```
(depvar1 indepvars1, equation(numlist))
        (depvar2 indepvars2, equation(numlist))
        [(depvar3 indepvars3, equation(numlist))]
        [...]
        [if] [in]
```

where **equation**(*numlist*) specifies equation numbers. Equation numbers are indexed 1,2,... up to the total number of equations (i.e. response variables) included in the model.

and the syntax of *random\_part* is

[ ... ] [level2(levelvar: [varlist] [, random\_part\_options])]
level1(levelvar: [varlist] [, random\_part\_options])



#### Examples

IMPORTANT. The following examples will only work on your computer once you have installed MLwiN and once you have told **runmlwin** the mlwin.exe file address. See **Remarks on installation instructions** above.

```
(a) Continuous response models
```

Two-level models

Setup . use http://www.bristol.ac.uk/cmm/media/runmlwin/tutorial, clear Two-level random-intercept model, analogous to xtreg. (See page 28 of the MLwin User Manual) (You will need to click the "Resume macro" button twice in MLwiN to fit the model.) . runmlwin normexam cons standlrt, level2(school: cons) level1(student: cons) Two-level random-intercept and random-slope (coefficient) model (See page 59 of the MLwin User Manual) . runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons) Refit the model, where this time we additionally calculate the level 2 residuals (See page 59 of the MLwin User Manual) . runm]win normexam cons stand]rt. level2 (school: cons stand]rt. residuals(u)) level1 (student: cons) Refit the model surpressing the two pauses in MLwiN (See page 59 of the MLwin User Manual) . runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons) nopause Two-level random-intercept and random-slope (coefficient) model with a complex level 1 variance function (See page 99 of the MLwin User Manual) matrix A = (1,1,0,0,0,1). runmlwin normexam cons standlrt girl,) level2(school: cons standlrt) level1(student: cons standlrt qirl, elements(A)) Multivariate response models Setup

. use http://www.bristol.ac.uk/cmm/media/runmlwin/gcsemv1, clear

Random-intercept bivariate response model (See page 214 of the MLwiN User Manual)

. runmlwin (written cons female, eq(1)) (csework cons female, eq(2)), level2(school: (cons, eq(1))
 (cons, eq(2))) level1(student: (cons, eq(1)) (cons, eq(2)))

Cross-classified models

🥹 Bristol University   Centr	e for Multilevel Modelling   runmlwin: Running MLwiN from wi	thin Stata - Mozilla Firefox		
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SOFTWARE	<u>University home</u> > <u>Centre for Multilevel Modelling</u> > <u>Software</u> > runr	mlwin		
MLwiN	Krunmlwin: Running MLwiN from within	Stata		
Realcom	runmlwin is a user written Stata command to fit mul	tilevel models in MI wiN from w	ithin Stata, Models car	be fit to both
MLPowSim	hierarchical and non-hierarchical (cross-classified ar	nd multiple membership) data	structures and to both	univariate
runmlwin	and multivariate responses. Models can be fit to con	tinuous, categorical (binary, pr	roportion, nominal, ord	inal) and
Presentations	count data.			
	The multilevel models fitted by <b>runmiwin</b> are analog	jous to those fitted by the Stat	a's xtmixed, xtmelog	it and
User Forum	xtmepoisson commands and by the user written gi	lamm command.		
CMM software support	running the <b>runmlwin</b> command in Stata carries out	: the following steps:		
	1. Writes an MLwiN macro for the specified multilev	vel model.		
	2. Opens MLwiN and runs the MLwiN macro.			
	<ol><li>Pauses MLwiN once the model is specified. This</li></ol>	allows the user to check that	the model is specified	as expected.
	4. Fits the model in MLwiN.			
	5. Pauses MLWIN once the model has been fitted (i	i.e. converged). This allows th	e user to examine the	model results.
	6. Stores and displays the model results in Stata			
	MLwiN and Stata are both required to use <b>runmlwin</b>	I.		
	Download			
	<b>runmlwin</b> is now available as a beta release. Users currently developing <b>runmlwin</b> so that jusers will so <u>Manual</u> .	should be able to fit all models on be able to additionally fit al	s in the <u>MLwiN User Ma</u> I models in the <u>MCMC</u>	anual. We are <u>MLwiN</u>

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MLwiN	KePresentations using runmlwin	
Realcom	<ul> <li>University of Bristol, Mplus/MlwiN User Group (MUGS) meeting (14th June 2011)</li> </ul>	
MLPowSim	<ul> <li><u>Slides</u> (PDF, 1.0mb)</li> </ul>	
runmlwin	<ul> <li>Stata do-file (do, 0.1mb) to replicate all analyses presented in the slides.</li> </ul>	
→ Presentations	- Madern Medeling Methods (M2) Conference, University of Connecticut (26th Mey 2011)	
Examples	Slides (PDF 3 0mb)	
User Forum	<ul> <li><u>State do-file</u> (do, 0.3mb) to replicate all analyses presented in the slides.</li> </ul>	
CMM software support		
	<ul> <li>2011 American Sociological Association Spring Methodology Conference, Tilburg University (20th Ma and a second second second second second second second second second second seco</li></ul>	iy 2011)
	<ul> <li><u>Slides</u> (PDF, 1.0mb)</li> <li>State de file (de, 0.1mb) to replicate all applyings presented in the slides.</li> </ul>	
	• <u>Stata do-file</u> (do, o, filip) to replicate all analyses presented in the slides.	
	<ul> <li>University of Bristol, e-Stat meeting (7th April 2011)</li> </ul>	
	<ul> <li><u>Slides</u> (PDF, 1.7mb)</li> </ul>	
	<ul> <li>Stata do-file (do, 0.1mb) to replicate all analyses presented in the slides.</li> </ul>	
	8th International Amsterdam Multilevel Conference (17th March 2011)	
	<ul> <li><u>Slides</u> (PDF, 2.0mb)</li> </ul>	
	<ul> <li>Stata do-file (do, 0.1mb) to replicate all analyses presented in the slides.</li> </ul>	

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ML with	🖗 Examples using runmlwin			
Beeleen				
Neaiconi M. BassiCire	MLwiN User Manual			
MLPOWSIM	These do-files and log files replicate the analyses report	ed in the <u>MLwiN User Mar</u>	nual.	
runmiwin		10 - ( #		and the set of the set
Presentations	We have also not vet attempted to replicate the analysis	or 19 of the manual as ho in Chanter 17	) models are fitted in th	iose chapters.
→ Examples				
User Forum	1 - Introducing Multilevel Models			
CMM software support	<ul> <li>2 - Introduction to Multilevel Modelling (do   log)</li> </ul>			
	<ul> <li>3 - Residuals (do   log)</li> </ul>			
	• 4 - Random Intercept and Random Slope Models (dd dd/dd/dd/dd/dd/dd/dd/dd/dd/dd/dd/dd/d	<u>e   log)</u>		
	• 5 - Graphical Procedures for Exploring the Model (directly should be added a straight directly sh	<u>o   log</u> )		
	<ul> <li>6 - Contextual Effects (<u>do</u>   <u>log</u>)</li> </ul>			
	• 7 - Modelling the Variance as a Function of Explanat	ory Variables ( <u>do</u>   <u>log</u> )		
	<ul> <li>8 - Getting Started with your Data</li> </ul>			
	9 - Logistic Models for Binary and Binomial Response	es ( <u>do   log</u> )		
	<ul> <li>10 - Multinomial Logistic Models for Unordered Cate</li> </ul>	gorical Responses ( <u>do</u>   <u>l</u>	<u>og</u> )	
	<ul> <li>11 - Fitting an Ordered Category Response Model (</li> </ul>	<u>do   log</u> )		
	<ul> <li>12 - Modelling Count Data (do   log)</li> </ul>			
	<ul> <li>13 - Fitting Models to Repeated Measures Data (do</li> </ul>	<u>log</u> )		
	<ul> <li>14 - Multivariate Response Models (<u>do</u>   <u>log</u>)</li> </ul>			-
	15 - Diagnostics for Multilevel Models (do   log)			

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Replicate the entire MLwiN User Manual using runmlwin     by GeorgeLeckie » Mon Apr 18, 2011 5:30 pm	0	68	by GeorgeLeckie 🖪 Mon Apr 18, 2011 5:30 pm	
Welcome to the runmlwin discussion forum by GeorgeLeckie » Fri Apr 01, 2011 4:06 pm	0	87	by GeorgeLeckie 🖪 Fri Apr 01, 2011 4:06 pm	
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TOPICS	REPLIES	VIEWS	LAST POST	1
TOPICS MCMC estimation by janna » Fri Apr 08, 2011 9:29 am	REPLIES	VIEWS 116	LAST POST by ChrisCharlton 🖪 Thu Jun 09, 2011 10:09 am	ĺ
TOPICS MCMC estimation by janna » Fri Apr 08, 2011 9:29 am error: too few quotes by laura » Wed May 18, 2011 5:54 pm	REPLIES 5 4	VIEWS 116 73	LAST POST by ChrisCharlton Thu Jun 09, 2011 10:09 am by laura Thu May 19, 2011 1:01 pm	
TOPICS         Image:	REPLIES   5   4   4	VIEWS 1116 73 145	LAST POST by ChrisCharlton C Thu Jun 09, 2011 10:09 am by laura C Thu May 19, 2011 1:01 pm by GeorgeLeckie C Thu May 12, 2011 3:51 pm	
TOPICS         Image: Stress of the second	REPLIES       5       4       4       1	VIEWS 1116 73 145 59	LAST POST by ChrisCharlton Thu Jun 09, 2011 10:09 am by laura Thu May 19, 2011 1:01 pm by GeorgeLeckie Thu May 12, 2011 3:51 pm by GeorgeLeckie Thu May 12, 2011 11:12 am	
TOPICS         Image: Stress of the stress	REPLIES         5         4         4         1         2	VIEWS 1116 73 145 59 107	LAST POST by ChrisCharlton  Thu Jun 09, 2011 10:09 am by laura Thu May 19, 2011 1:01 pm by GeorgeLeckie Thu May 12, 2011 3:51 pm by GeorgeLeckie Thu May 12, 2011 11:12 am by ChrisCharlton Wed Apr 27, 2011 11:06 pm	
TOPICS         Image: Stress process of the second pr	REPLIES         5         4         4         1         2         3	VIEWS 1116 73 145 59 107 86	LAST POST by ChrisCharlton  Thu Jun 09, 2011 10:09 am by laura  Thu May 19, 2011 1:01 pm by GeorgeLeckie  Thu May 12, 2011 3:51 pm by GeorgeLeckie  Thu May 12, 2011 11:12 am by ChrisCharlton  Wed Apr 27, 2011 11:06 pm by GeorgeLeckie  Fri Apr 22, 2011 10:56 am	

# Citing runmlwin

- If you use runmlwin in your work, please cite runmlwin
- Leckie, G. and Charlton, C. (2011) *runmlwin: Stata module for fitting multilevel models in the MLwiN software package*. Centre for Multilevel Modelling, University of Bristol.
- We can then add you to the list of papers using runmlwin on our website
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