

Social disadvantage and infant mortality: the birth weight paradox revisited

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First observed by Yerushalmy (1964, 1971) and interpreted as BW modifying the effect of many factors associated with infant mortality:

BW paradox

- **Smoking** known risk factor for low BW.
- **Low BW** babies born to smokers lower mortality than those of non-smokers:

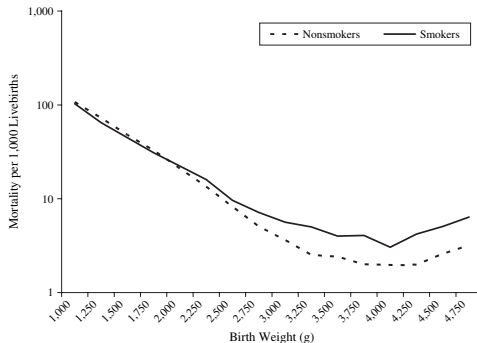
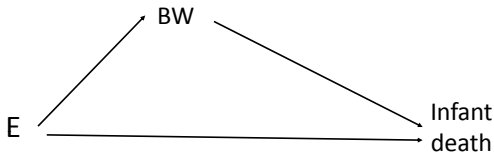


Figure: Birth-weight-specific infant mortality curves, US, 1991 (Hernandez-Diaz, AJE 2006)

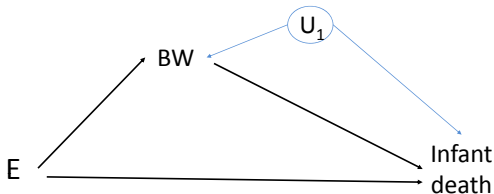
The low birth weight paradox: collider bias?

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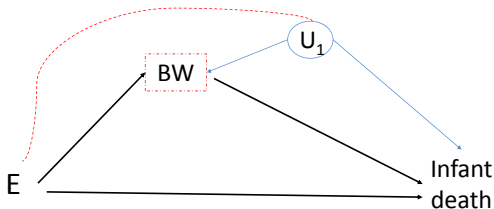
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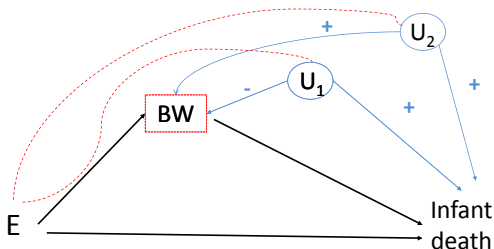
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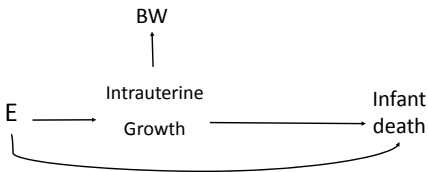
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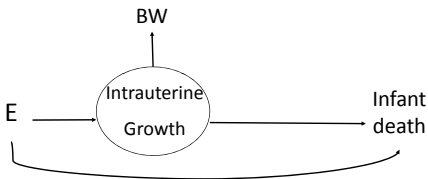
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- neither intrauterine dimensions are usually available in large observational studies.
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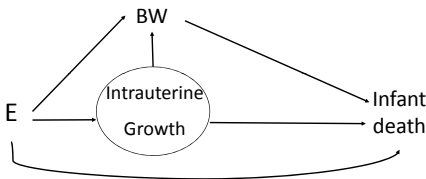
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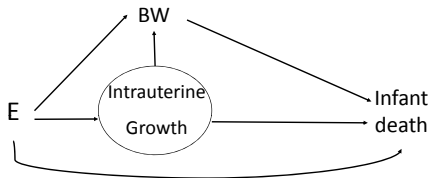
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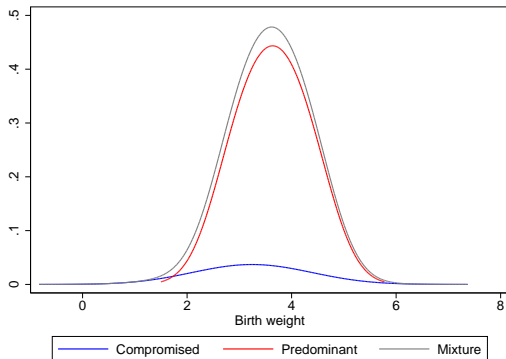
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But how can we proceed without information on intrauterine growth?



Wilcox (1983,2001) suggested that there are two sub-populations of newborns:

- (a) **predominant**: mostly term babies,
- (b) **compromised**: mostly pre-term babies and small-for-gestational-age.

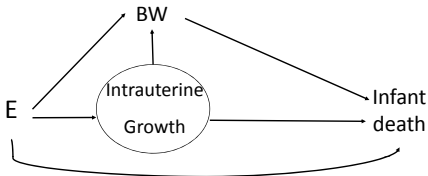


- The model can be reformulated in terms of these classes.

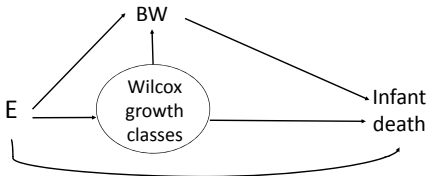
- Assuming that the birth weight distribution for each sub-population is normal,

- and that the population is well-mixed,

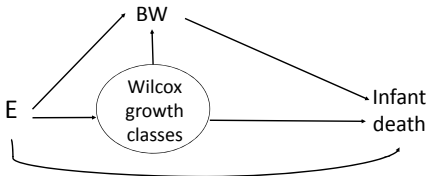
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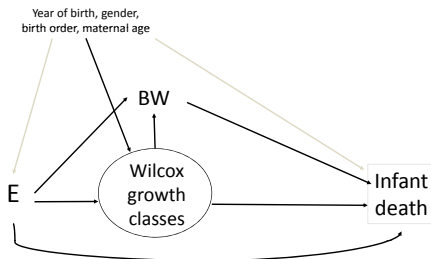
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With this more general theoretical framework, we reconsider the two main questions.

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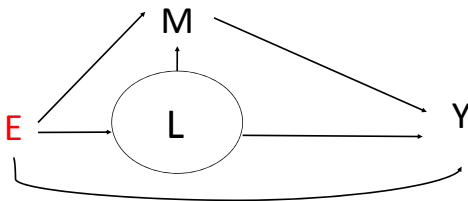
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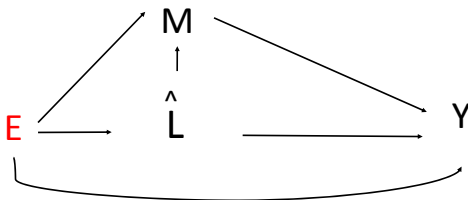
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- BW: potential mediator (M); “Disadvantage”: exposure (E); Infant mortality: outcome (Y); “Intrauterine growth”: intermediate confounder (L).
- Replacing L with $\hat{L} = \Pr(L = 1)$ (1: compromised, 0: predominant),

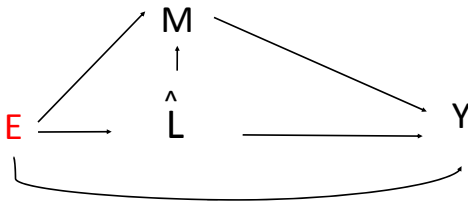


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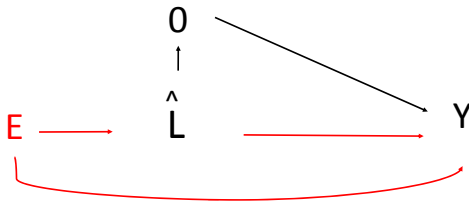
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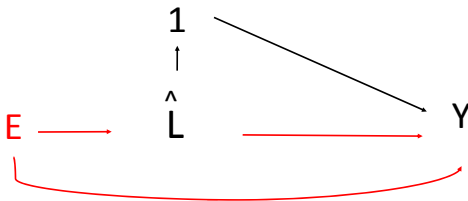
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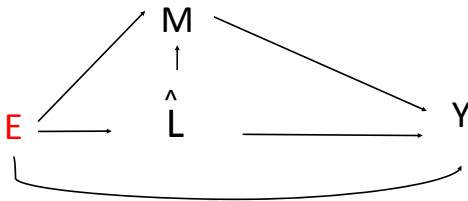
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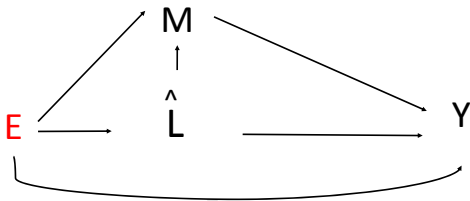
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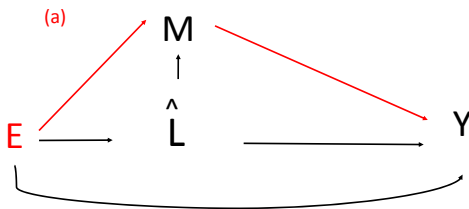
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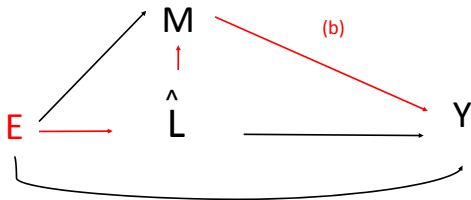
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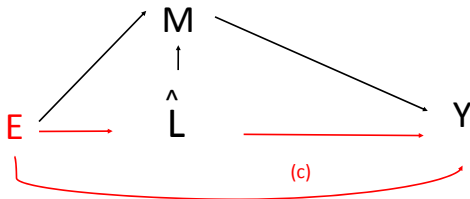
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Estimands (CDE(m) and PNDE, TNIE) are expressed as OR contrasts.

Assumptions:

No interference, consistency, conditional exchangeability, and, because of L , either:

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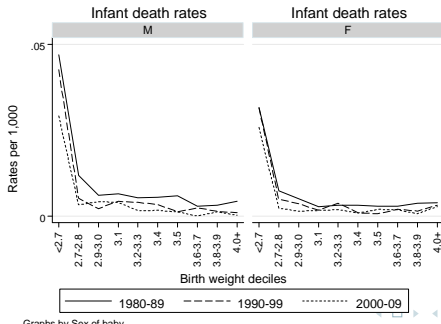
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Natural direct and indirect effects of low maternal education

VERY PRELIMINARY RESULTS- SEs not yet corrected

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	ln OR	(SE)	ln OR	(SE)
CDE(0)	–	–	0.205	(0.076)
CDE(1)	–	–	0.206	(0.076)
PNDE	0.221	(0.082)	0.227	(0.077)
TNIE	0.011	(0.007)	-0.012	(0.005)
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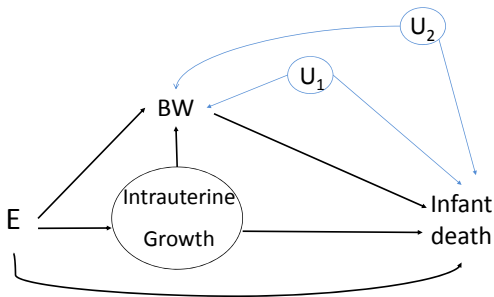
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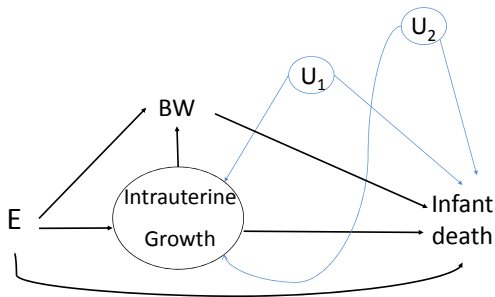
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- Results depends on **strong and partly unverifiable assumptions**, although similarity of results from alternative parametric specifications are reassuring.
- **Estimation** of mediation effects and their SEs raises several problems. There are issues with:
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The authors alone are responsible for the interpretation of the data.

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- The total causal effect (TCE):

$$TCE^{OR} = \frac{E[Y(1)]/\{1 - E[Y(1)]\}}{E[Y(0)]/\{1 - E[Y(0)]\}}$$

- The natural direct effect (NDE):

$$NDE^{OR} = \frac{E[Y(1, M(0))]/\{1 - E[Y(1, M(0))]\}}{E[Y(0, M(0))]/\{1 - E[Y(0, M(0))]\}}$$

- The natural indirect effect (NIE):

$$NIE^{OR} = \frac{E[Y(1, M(1))]/\{1 - E[Y(1, M(1))]\}}{E[Y(1, M(0))]/\{1 - E[Y(1, M(0))]\}}$$

where $Y(x)$ is the potential value of Y that would have occurred had X been set to x and $Y(x, m)$ the potential value of Y that would have occurred had X been set to x and M to m

Mat Education	Birth weight ≥ 2.5 kg		Birth weight < 2.5 kg	
	Low	High	Low	High
Births	92,704	59,141	4,393	4,128
Deaths	220	222	225	195
Rates (x 1,000)	2.4	3.8	51.24	47.2
Sex-adjusted OR <i>heterog test (p)</i>	1.58 (1.31, 1.91)		0.92 (0.76, 1.12)	
		<i>(0.031)</i>		
<i>Adjusted² OR heterog test (p)</i>	1.23 (1.01, 1.49)		0.92 (0.76, 1.12)	
		<i>(0.036)</i>		

	Variable	Class 1	Class 2
For μ	Intercept	3.51	3.65
	sex	-	-
	year birth	-	+
	mat age	+	+
	birth order	-	+
For σ	Intercept	0.90	0.45
For π	sex	-	
	Mat educ	+	

About 10% of births predicted to be “compromised”.

- There is another source of bias: conditioning on live birth.
- Still births are a form of competing event, reducing the denominator of possible infant deaths.
- Consider the **composite outcome** of Infant death or Still birth
(Kramer *et al.* , 2014):

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PNDE	0.221	(0.082)	0.174	(0.067)
TNIE	0.011	(0.007)	0.018	(0.008)
TCE	0.232	(0.082)	0.192	(0.066)